

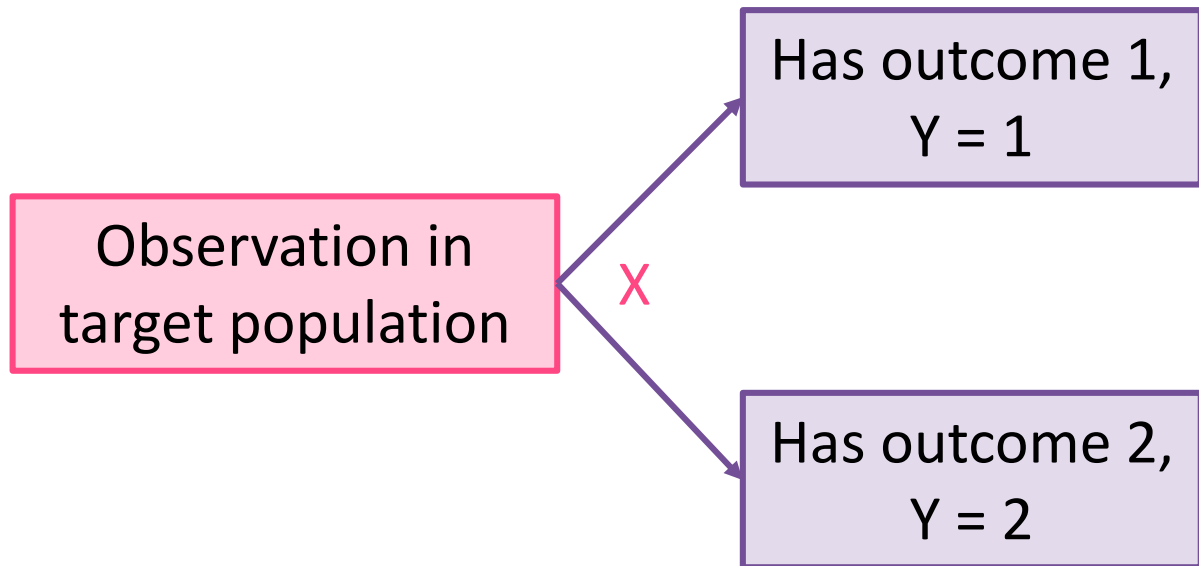
Statistical inference for association studies in the presence of binary outcome misclassification

Kimberly A. Hochstedler and Martin T. Wells

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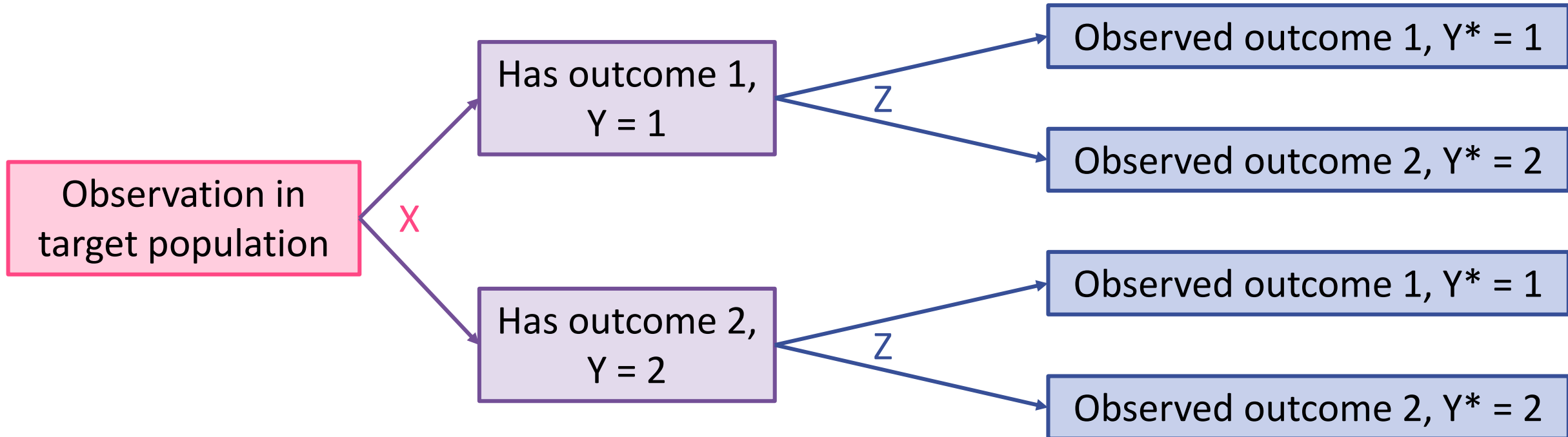
Problem setting

- Interested in the association between X and the **binary variable** Y .



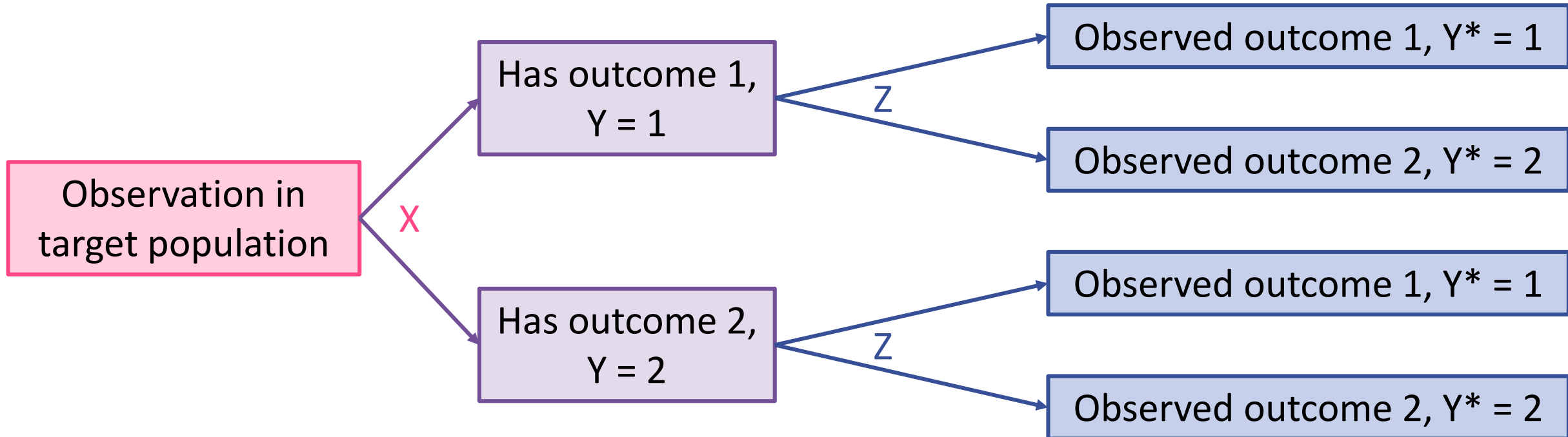
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- Measure Y using an instrument that is **not always accurate**, and obtain Y^* .



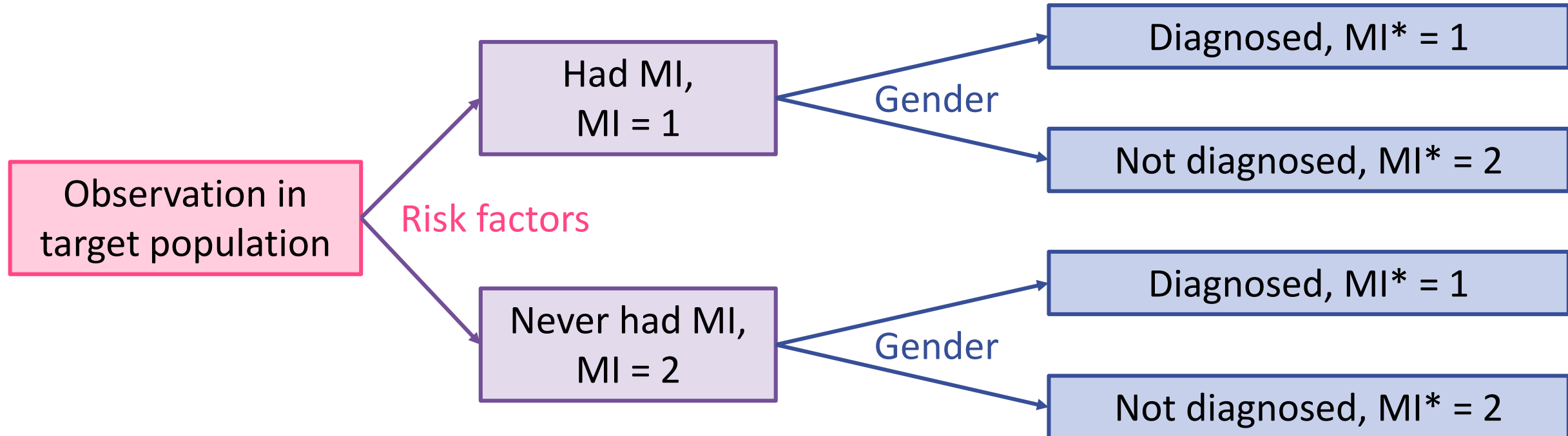
Problem setting

- Interested in the association between X and the **binary variable** Y .
- Measure Y using an instrument that is **not always accurate**, and obtain Y^* .
- A third variable, Z , is related to the **misclassification mechanism**.

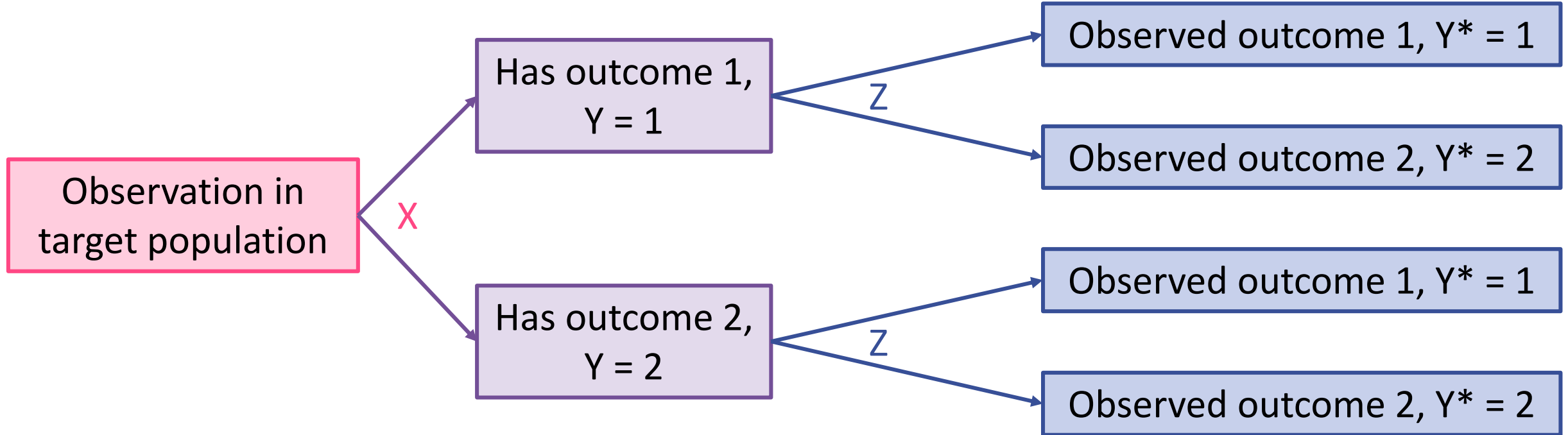


Example

- Interested in the association between **Risk Factors** and the **binary variable MI**.
- Measure **MI** using **self-reported medical diagnoses**, and obtain **MI***.
- A third variable, **gender**, is related to the **misclassification mechanism**.



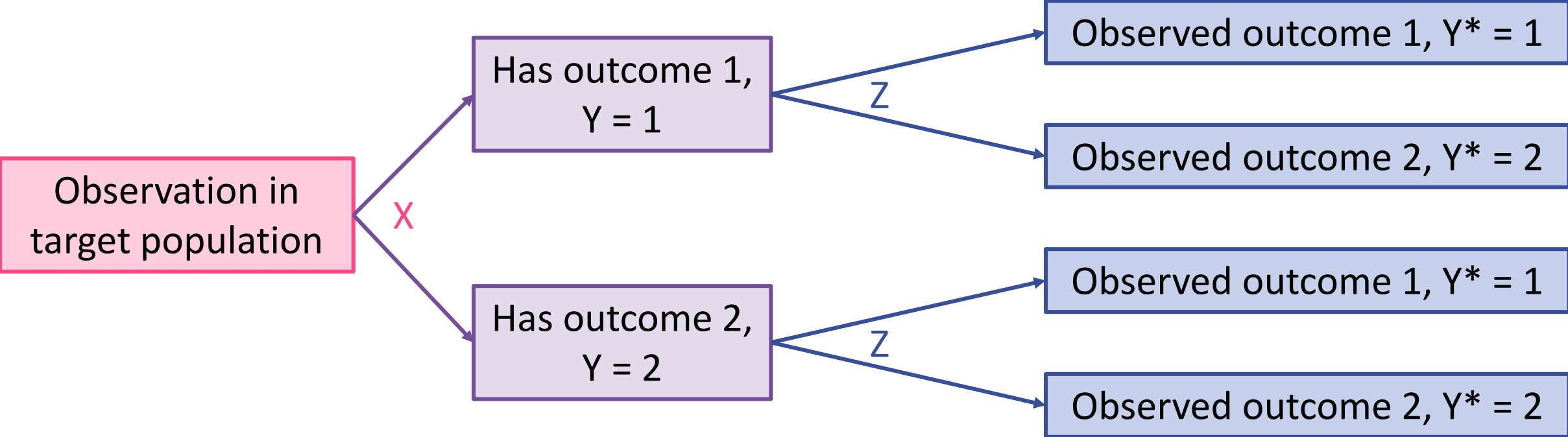
Misclassification model



True outcome mechanism: $\text{logit}\{P(Y = j|X; \beta)\} = \beta_{j0} + \beta_{jX}X$

Observation mechanism: $\text{logit}\{P(Y^* = k|Y = j, Z; \gamma)\} = \gamma_{kj0} + \gamma_{kjZ}Z$

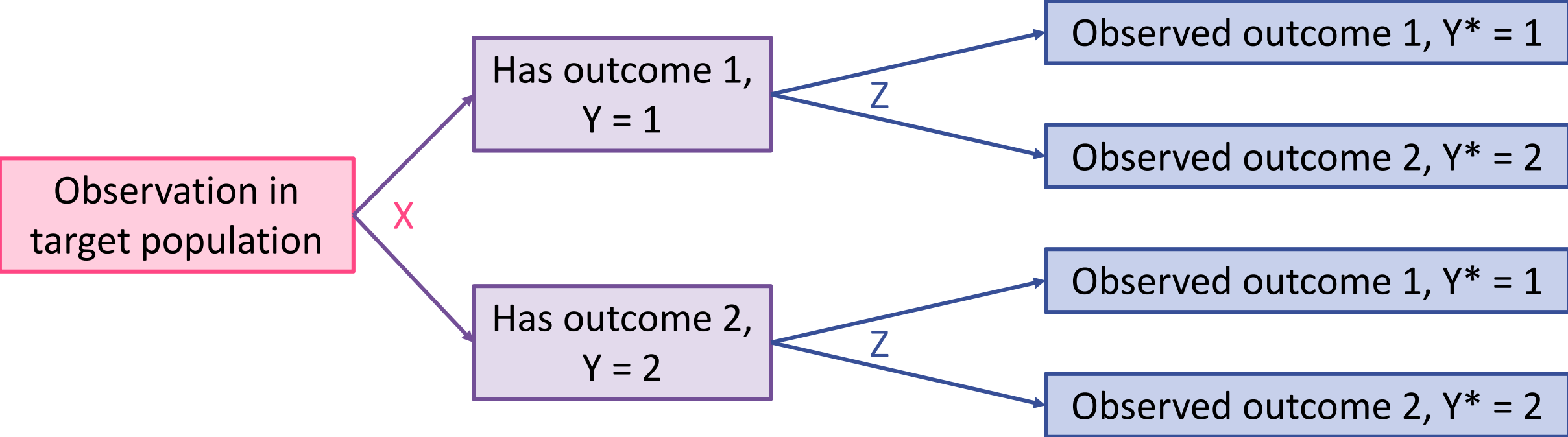
Misclassification model



$$P(Y_i = j | X; \beta) = \pi_{ij} = \frac{\exp\{\beta_{j0} + \beta_{jX} X_i\}}{1 + \exp\{\beta_{j0} + \beta_{jX} X_i\}}$$

$$P(Y_i^* = k | Y_i = j, Z; \gamma) = \pi_{ikj}^* = \frac{\exp\{\gamma_{kj0} + \gamma_{kjZ} Z_i\}}{1 + \exp\{\gamma_{kj0} + \gamma_{kjZ} Z_i\}}$$

Misclassification model

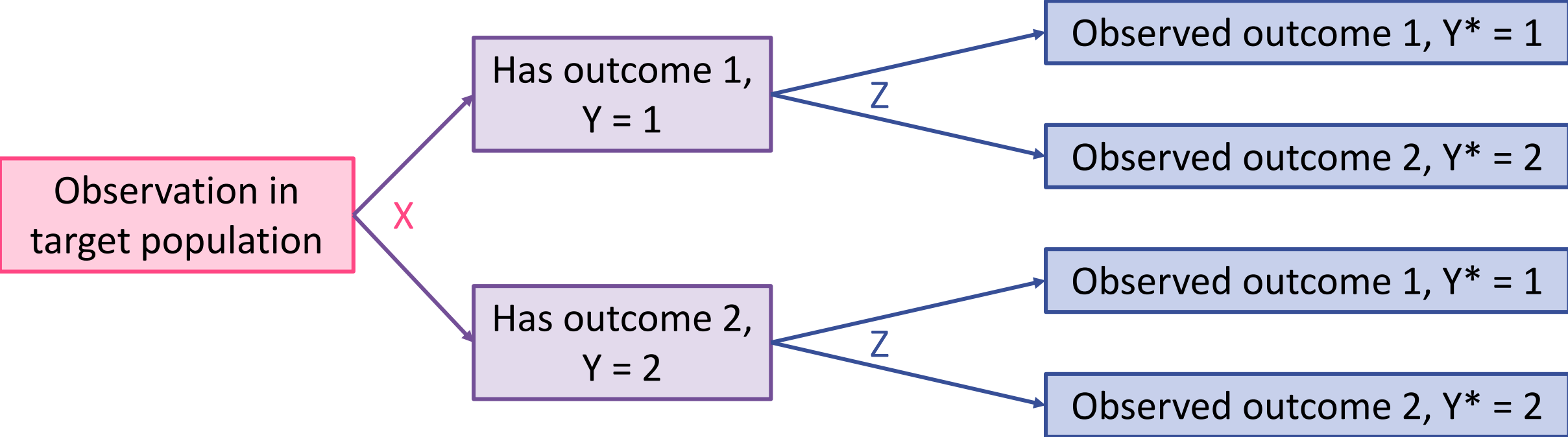


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Primary interest: Estimating β

$$P(Y_i^* = k | Y_i = j, Z; \gamma) = \pi_{ikj}^* = \frac{\exp\{\gamma_{kj0} + \gamma_{kjZ} Z_i\}}{1 + \exp\{\gamma_{kj0} + \gamma_{kjZ} Z_i\}}$$

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Secondary interest: Estimating γ

Complete data log-likelihood

- Y (true outcome) is a latent variable, but let's pretend we know it:

$$\ell_{complete}(\beta, \gamma; X, Z) = \sum_{i=1}^N \left[\sum_{j=1}^2 y_{ij} \log\{P(Y_i = j | X_i)\} + \sum_{j=1}^2 \sum_{k=1}^2 y_{ij} y_{ik}^* \log\{P(Y_i^* = k | Y_i = j, Z_i)\} \right]$$

Complete data log-likelihood

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True outcome portion

Observed outcome, given true outcome portion

The diagram illustrates the decomposition of the complete data log-likelihood function. The first term, $\sum_{j=1}^2 y_{ij} \log\{P(Y_i = j | X_i)\}$, is identified as the 'True outcome portion' and is defined as $y_{ij} = \mathbb{I}\{Y_i = j\}$. The second term, $\sum_{j=1}^2 \sum_{k=1}^2 y_{ij} y_{ik}^* \log\{P(Y_i^* = k | Y_i = j, Z_i)\}$, is identified as the 'Observed outcome, given true outcome portion'.

Complete data log-likelihood

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$y_{ij} = \mathbb{I}\{Y_i = j\}$

True outcome portion

Observed outcome, given true outcome portion

$$= \sum_{i=1}^N \left[\sum_{j=1}^2 y_{ij} \log\{\pi_{ij}\} + \sum_{j=1}^2 \sum_{k=1}^2 y_{ij} y_{ik}^* \log\{\pi_{ikj}^*\} \right]$$

Complete data log-likelihood

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Estimation with the EM Algorithm



Estimation with the EM Algorithm

Expectation Step

Maximization Step



$$w_{ij} = P(Y_i = j | Y_i^*, X, Z) = \sum_{k=1}^2 \frac{y_{ik}^* \pi_{ikj}^* \pi_{ij}}{\sum_{l=1}^2 \pi_{ikl}^* \pi_{il}}$$

Estimation with the EM Algorithm

Expectation Step

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$$Q = \sum_{i=1}^N \left[\sum_{j=1}^2 w_{ij} \log\{\pi_{ij}\} + \sum_{j=1}^2 \sum_{k=1}^2 w_{ij} y_{ik}^* \log\{\pi_{ikj}^*\} \right]$$

Estimation with the EM Algorithm

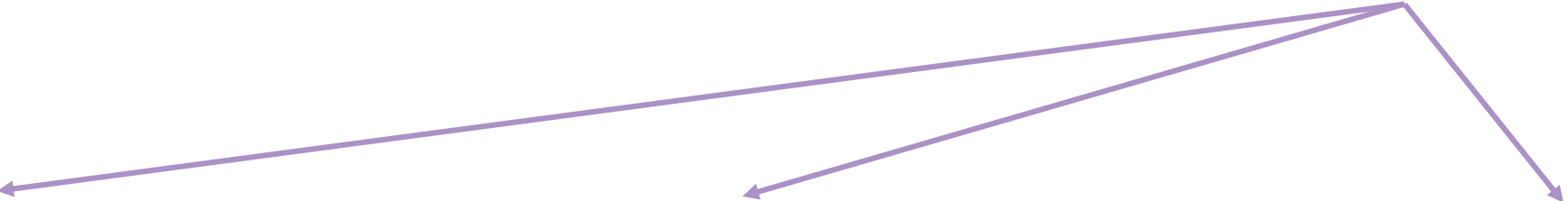
Expectation Step

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$$Q_{\beta} = \sum_{i=1}^N \left[\sum_{j=1}^2 w_{ij} \log\{\pi_{ij}\} \right]$$

$$Q_{\gamma_{k1}} = \sum_{i=1}^N \left[\sum_{k=1}^2 w_{i1} y_{ik}^* \log\{\pi_{ik1}^*\} \right]$$

$$Q_{\gamma_{k2}} = \sum_{i=1}^N \left[\sum_{k=1}^2 w_{i2} y_{ik}^* \log\{\pi_{ik2}^*\} \right]$$

Label switching

- **Label switching:** When a mixture model likelihood is **invariant under relabeling** of the mixture components, resulting in **multimodal likelihood functions**.

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$$\begin{aligned} \ell_{complete}(\beta, \gamma; X, Z) &= \sum_{i=1}^N \left[y_{i1} \log\{\pi_{i1}\} + y_{i2} \log\{\pi_{i2}\} \right. \\ &\quad \left. + y_{i1}y_{i1}^* \log\{\pi_{i11}^*\} + y_{i1}y_{i2}^* \log\{\pi_{i21}^*\} + y_{i2}y_{i1}^* \log\{\pi_{i12}^*\} + y_{i2}y_{i2}^* \log\{\pi_{i22}^*\} \right] \\ \ell_{complete}(\beta, \gamma; X, Z) &= \sum_{i=1}^N \left[y_{i2} \log\{\pi_{i2}\} + y_{i1} \log\{\pi_{i1}\} \right. \\ &\quad \left. + y_{i2}y_{i1}^* \log\{\pi_{i12}^*\} + y_{i2}y_{i2}^* \log\{\pi_{i22}^*\} + y_{i1}y_{i1}^* \log\{\pi_{i11}^*\} + y_{i1}y_{i2}^* \log\{\pi_{i21}^*\} \right] \end{aligned}$$

Label switching

- Suppose we have a single predictor X and a single predictor Z:

$$\begin{aligned} & \sum_{i=1}^N \left[y_{i1} \beta_0 + y_{i1} x_i \beta_X - (y_{i1} + y_{i2}) \log\{1 + \exp\{\beta_0 + x_i \beta_X\}\} \right. \\ & \quad + y_{i1} y_{i1}^* \gamma_{110} + y_{i1} y_{i1}^* z_i \gamma_{11Z} - (y_{i1}^* + y_{i2}^*) y_{i1} \log\{1 + \exp\{\gamma_{110} + z_i \gamma_{11Z}\}\} \\ & \quad \left. + y_{i2} y_{i1}^* \gamma_{120} + y_{i2} y_{i1}^* z_i \gamma_{12Z} - (y_{i1}^* + y_{i2}^*) y_{i2} \log\{1 + \exp\{\gamma_{120} + z_i \gamma_{12Z}\}\} \right] \\ = & \sum_{i=1}^N \left[y_{i2} (-\beta_0) + y_{i2} x_i (-\beta_X) - (y_{i1} + y_{i2}) \log\{1 + \exp\{-\beta_0 + x_i (-\beta_X)\}\} \right. \\ & \quad + y_{i2} y_{i1}^* \gamma_{120} + y_{i2} y_{i1}^* z_i \gamma_{12Z} - (y_{i1}^* + y_{i2}^*) y_{i2} \log\{1 + \exp\{\gamma_{120} + z_i \gamma_{12Z}\}\} \\ & \quad \left. + y_{i1} y_{i1}^* \gamma_{110} + y_{i1} y_{i1}^* z_i \gamma_{11Z} - (y_{i1}^* + y_{i2}^*) y_{i1} \log\{1 + \exp\{\gamma_{110} + z_i \gamma_{11Z}\}\} \right] \end{aligned}$$

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Label switching

- There are two sets of parameters that yield the **exact same likelihood value**.

$$\beta_0, \beta_X, \gamma_{110}, \gamma_{11Z}, \gamma_{120}, \gamma_{12Z}$$

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$$-\beta_0, -\beta_X, \gamma_{120}, \gamma_{12Z}, \gamma_{110}, \gamma_{11Z}$$

Correcting label switching

- There is a quantity that has different values when each parameter set is used to compute it:

$$P(Y_i^* = k | Y_i = j, Z; \gamma) = \pi_{ikj}^* = \frac{\exp\{\gamma_{kj0} + \gamma_{kjZ} Z_i\}}{1 + \exp\{\gamma_{kj0} + \gamma_{kjZ} Z_i\}}$$

Correcting label switching

Algorithm 1 Correcting label switching in binary outcome misclassification models

Compute average π_{jj}^* for all $j \in \{1, 2\}$ using $\hat{\beta}$ and $\hat{\gamma}$.

if $\pi_{jj}^* > 0.50$ for all $j \in \{1, 2\}$ **then**

$$\hat{\beta}_{corrected} \leftarrow \hat{\beta}$$

$$\hat{\gamma}_{corrected} \leftarrow \hat{\gamma}$$

else

$$\hat{\beta}_{corrected} \leftarrow -\hat{\beta}$$

$$\hat{\gamma}_{corrected,k1} \leftarrow \hat{\gamma}_{k2}$$

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end if

Correcting label switching

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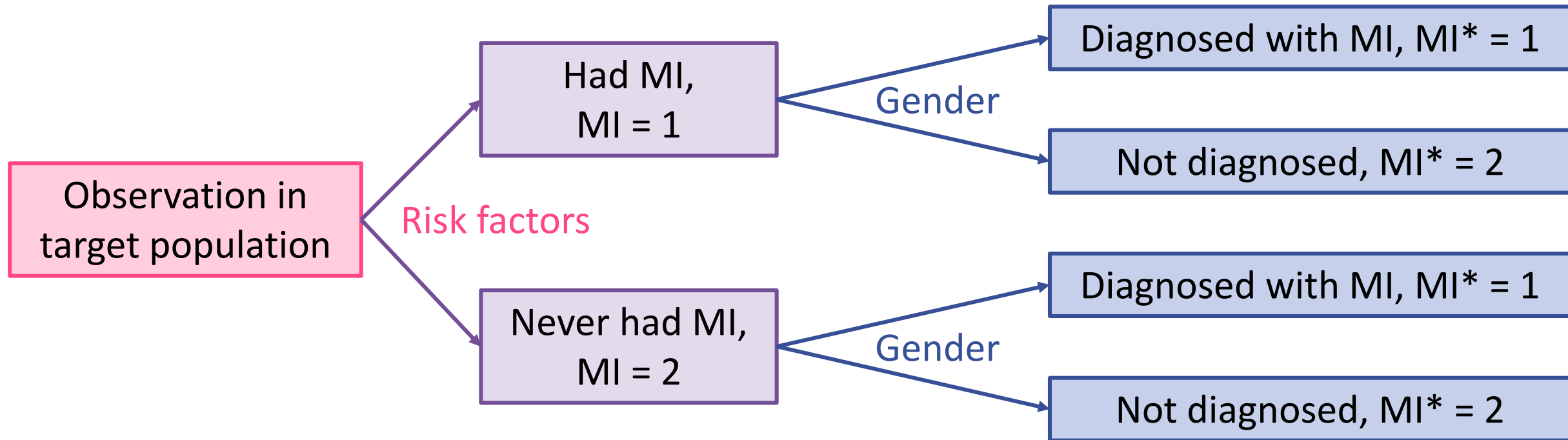
end if



Assumption:
Outcome categories are
correctly classified at
least 50% of the time.

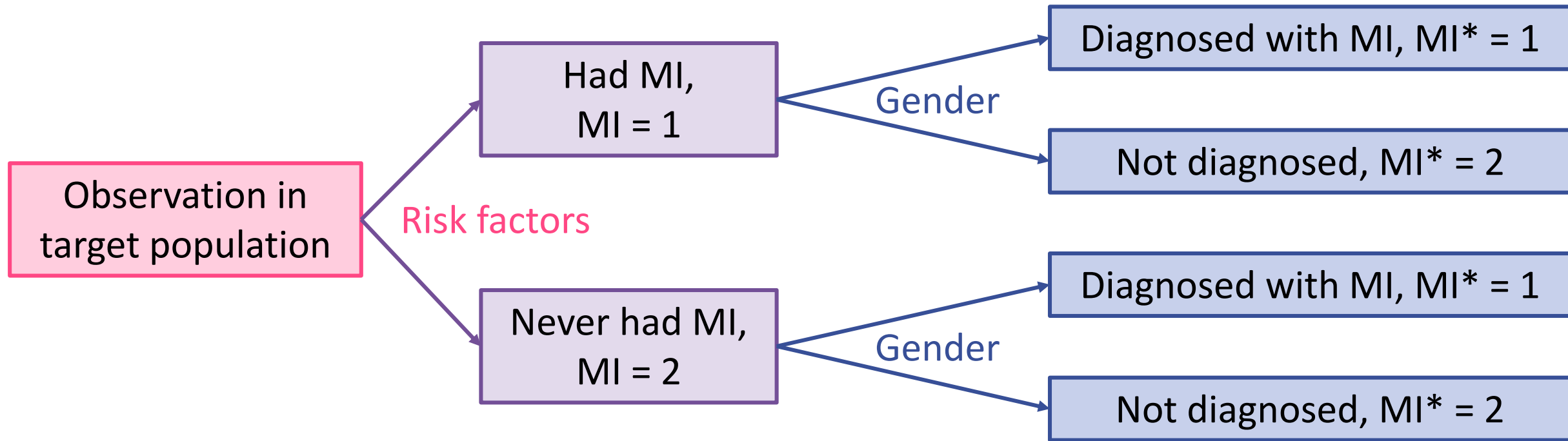
- Apply to EM estimates.

Applied Example



- **Goal:** Understand the risk factors for MI.
 - MI is suspected to be misdiagnosed differentially based on patient age and gender.
 - Data from 2020 MEPS survey.

Applied Example



- **Model for true MI:** $MI \sim \text{Smoking Status} + \text{Exercise Habits} + \text{Age}$
- **Model observed MI given true MI:** $MI^* \mid MI \sim \text{Age} + \text{Gender}$

Applied Example

- **Model for true MI:** $MI \sim \text{Smoking Status} + \text{Exercise Habits} + \text{Age}$
- **Model observed MI given true MI:** $MI^* \mid MI \sim \text{Age} + \text{Gender}$

	EM		Naive Analysis	
	Est.	SE	Est.	SE
β_0	-4.374	0.065	-3.576	0.078
β_{smoke}	1.544	0.107	0.635	0.109
$\beta_{exercise}$	0.303	0.126	0.184	0.084
β_{age}	0.094	0.010	0.059	0.003
γ_{110}	2.969	0.100	-	-
$\gamma_{11,gender}$	-1.766	0.036	-	-
$\gamma_{11,age}$	-0.198	0.005	-	-
γ_{120}	-3.580	0.112	-	-
$\gamma_{12,gender}$	-0.818	0.108	-	-
$\gamma_{12,age}$	0.084	0.005	-	-

Effects are attenuated when we do not account for misclassification of MI

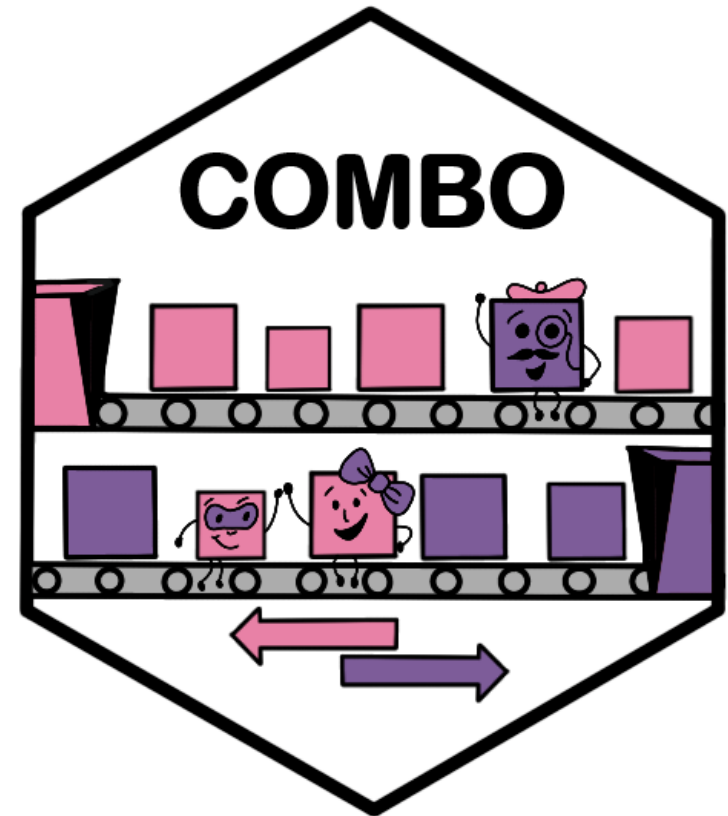
Applied Example

- **Model for true MI:** $MI \sim \text{Smoking Status} + \text{Exercise Habits} + \text{Age}$
- **Model observed MI given true MI:** $MI^* \mid MI \sim \text{Age} + \text{Gender}$

	Estimated Specificity $P(\text{no MI}^* \mid \text{no MI})$	Estimated Sensitivity $P(\text{MI}^* \mid \text{MI})$
Men	94.4%	76.3%
Women	97.1%	59.1%

Want to use this method?

- You're in luck!
- *COMBO* R Package (coming soon to a CRAN repository near you).
 - **Correcting Misclassified Binary Outcomes**



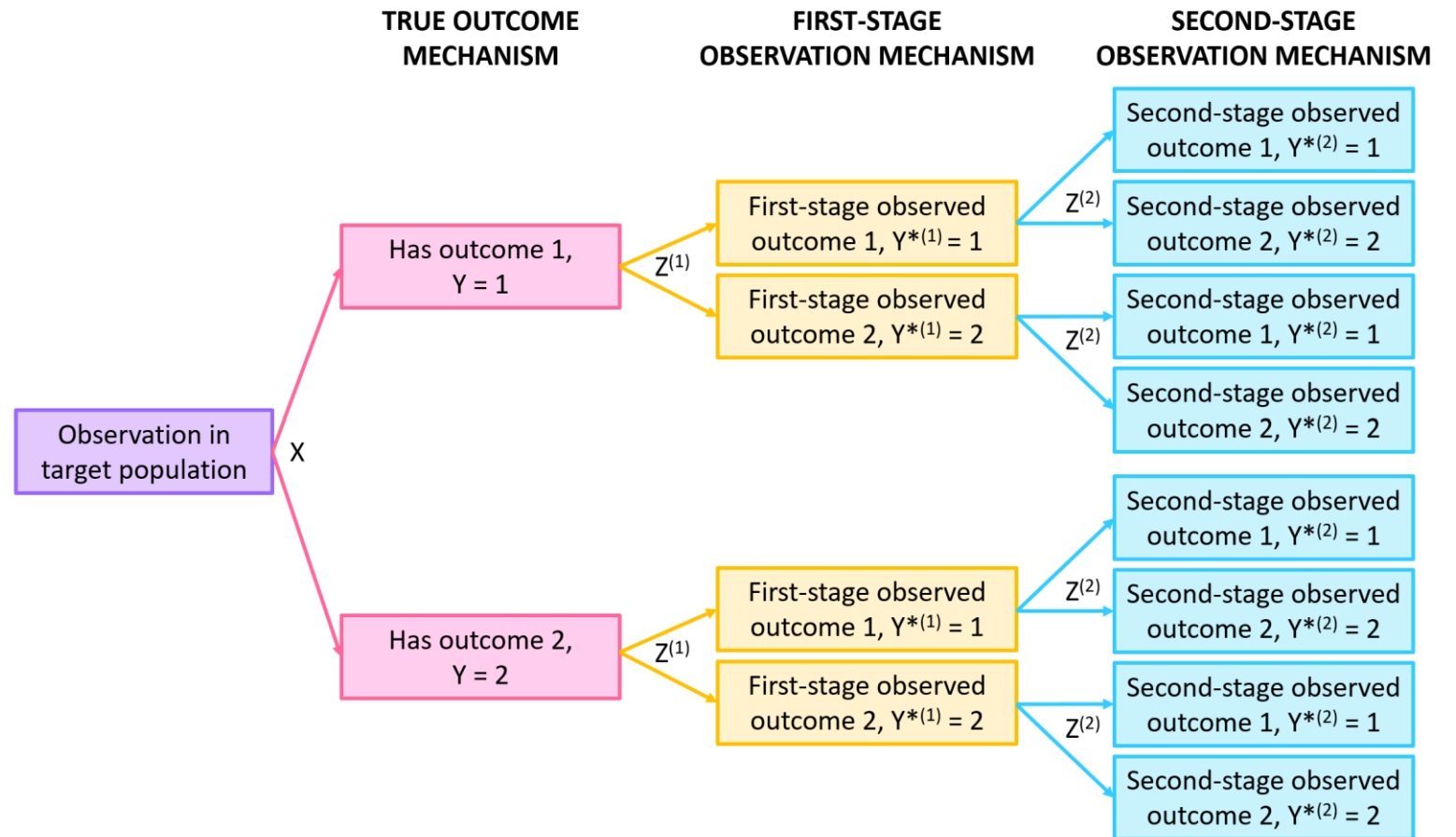
Conclusions and Next Steps

- We can use the proposed EM algorithm to **estimate associations** when a **binary outcome is potentially misclassified**.

Conclusions and Next Steps

- **Extensions:**

- Outcomes with more than 2 categories.
- More than one outcome “stage”.



Thank you!

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Notation cheat sheet and more info on
“COMBO” available at: bit.ly/R_COMBO